



UPPSALA
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HOW ARE WE RELATED TO EACH OTHER?

SIMPLE MODELS OF OUR GENETIC ANCESTRY

DOCENTSHIP LECTURE BY RAAZESH SAINUDIIN
DEPARTMENT OF MATHEMATICS, UPPSALA UNIVERSITY

2022-10-07



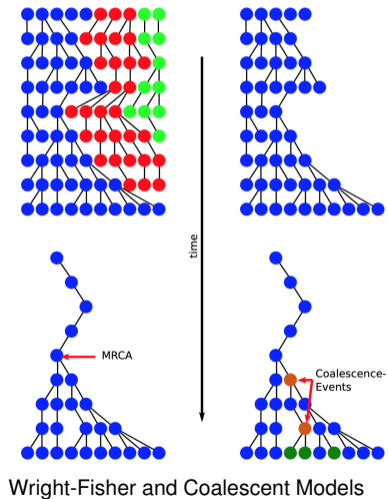
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- Introduce the Wright-Fisher model.
- Genealogies in the Wright-Fisher model.
- How to obtain the coalescent from the Wright-Fisher model.
- Properties of coalescent models



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Wright-Fisher and Coalescent Models



THE WRIGHT-FISHER MODEL: INTRODUCTION

- There are N individuals in each generation
- Sample current generation with replacement to produce next generation
- Repeat the above step every generation



Sewall Wright
(1889-1988)



Sir Ronald Fisher
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THE WRIGHT-FISHER MODEL: INTRODUCTION

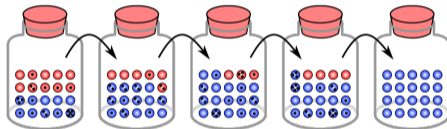
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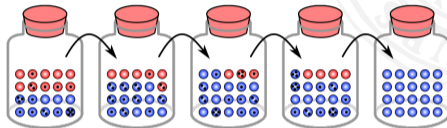
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Theorem (Offspring Distribution)

This reproduction mechanism of the W-F Model gives the number of offspring in the next generation for each individual in the current generation labelled by $1, 2, \dots, N$ as the Multinomial($N, (1/N, 1/N, \dots, 1/N)$) distribution .

Mathematical Preliminaries 0.1 – basic counting

Lemma (number of permutations of n elements)

The number of different ways of arranging n distinct elements into a sequence, a.k.a., the permutations of n , is $n! := n \cdot (n-1) \cdot (n-2) \cdots 2 \cdot 1$.

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For example, there are $3! = 3 \cdot 2 \cdot 1 = 6$ permutations of $\{1, 2, 3\}$. They are:

$(1, 2, 3), (2, 1, 3), (3, 1, 2), (1, 3, 2), (2, 3, 1), (3, 2, 1)$.

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Proof.

The proof follows from the the following construction:

- there are n possible choices for the first element in the sequence
- for each such choice there are $(n-1)$ elements left to choose from for the second element
- and so on, until there is exactly one element left for the n -th element in the sequence
- the total number of permutations follow from multiplying the number of possible choices in each of the n steps above giving $n \cdot (n-1) \cdot (n-2) \cdots 2 \cdot 1$.
- Note that there is only one way of arranging $\{\}$, an empty set without any element, so $0! = 1$.

□

Mathematical Preliminaries 0.2 – basic counting

Lemma (number of subsets of size k from a set with n elements)

The number of the k -element combinations, a.k.a., subsets of k elements from a set with n elements, is

$$\binom{n}{k} := \frac{n!}{k!(n-k)!} = \frac{n \cdot (n-1) \cdot (n-2) \cdots (n-k+1)}{k \cdot (k-1) \cdots 2 \cdot 1}$$

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For example, there are $\frac{5 \cdot 4 \cdot 3}{3 \cdot 2 \cdot 1} = \frac{60}{6} = 10$ 3-element combinations of $\{1, 2, 3, 4, 5\}$. They are:

$\{1, 2, 3\}, \{1, 2, 4\}, \{1, 2, 5\}, \{1, 3, 4\}, \{1, 3, 5\}, \{1, 4, 5\}, \{2, 3, 4\}, \{2, 3, 5\}, \{2, 4, 5\}, \{3, 4, 5\}$.

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Proof.

Follows by construction where,

- the numerator counts the number of ways to choose the first k elements from the set of n elements and arrange them in a row with regard to order
- the denominator disregards order as each k -element combination appears exactly $k!$ times

□



THE WRIGHT-FISHER MODEL: MULTINOMIAL OFFSPRING

- Label individuals in (parent) generation g_1 by 1, 2, 3, 4, 5 in the W-F Model with $N = 5$





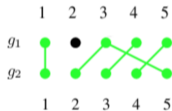
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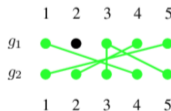
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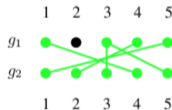


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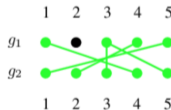
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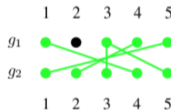
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- Probability of any of the N^N many offspring-parent relations is $\left(\frac{1}{N}\right)^N$
- Let X_i be offspring number of parent i .
What is $\Pr(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N)$?



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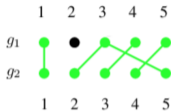


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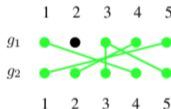
$$\Pr(X_1 = 1, X_2 = 0, X_3 = 2, X_4 = 1, X_5 = 1) \geq 2 \left(\frac{1}{5}\right)^5$$

But what is it exactly?

THE WRIGHT-FISHER MODEL: MULTINOMIAL OFFSPRING



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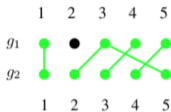
$$\Pr(X_1 = 1, X_2 = 0, X_3 = 2, X_4 = 1, X_5 = 1) = ?? \left(\frac{1}{5}\right)^5$$

Lemma (Multinomial Coefficient)

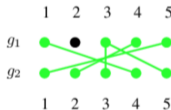
Let $\binom{N}{x_1, x_2, \dots, x_N}$ be the number of ways parent i chooses x_i many offspring, $\forall i = 1, 2, \dots, N$. Then

$$\binom{N}{x_1, x_2, \dots, x_N} = \frac{N!}{x_1! x_2! \cdots x_N!}$$

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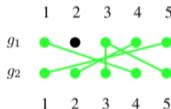
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- Before proving the Lemma let's apply it to our example

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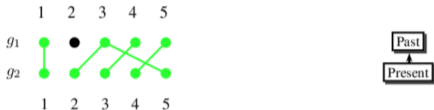
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THE WRIGHT-FISHER MODEL: MULTINOMIAL OFFSPRING



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- Before proving the Lemma let's apply it to our example
- $\binom{5}{1,0,2,1,1} = \frac{5!}{1!0!2!1!1!} = 60$
- If the lemma were true, then there are 60 possible distinct parent-offspring relations such that parent 1 has 1 offspring, parent 2 has none, parent 3 has 2 offspring, and parent 4 has 1 offspring and parent 5 has 1 offspring.

THE WRIGHT-FISHER MODEL: MULTINOMIAL OFFSPRING

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Proof.

The proof follows from the following construction:



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- From the remaining $N - x_1 - x_2$ individuals choose x_3 to be labelled as offspring of individual 3. Again, this can be done $\binom{N-x_1-x_2}{x_3}$ ways. [Proof continued...]



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Proof.

The proof continues from the above construction:

- Multiplying the number of choices at each step results in:

$$\binom{N}{x_1} \binom{N-x_1}{x_2} \binom{N-x_1-x_2}{x_3} \cdots = \frac{N!}{(N-x_1)!} \frac{(N-x_1)!}{(N-x_1-x_2)! x_2!} \frac{(N-x_1-x_2)!}{(N-x_1-x_2-x_3)! x_3!} \cdots$$

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- Cancellation results in the desired formula:

$$\frac{N!}{x_1! x_2! \cdots x_N!}$$



THE WRIGHT-FISHER MODEL: MULTINOMIAL OFFSPRING

Combining the last two observations and lemmas, we have proved the following theorem.

Theorem (Offspring Distribution of W-F Model is Multinomial)

Let X_i be the number of offspring of parent i , $\forall i = 1, 2, \dots, N$. Then,

$$\Pr(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N) = \binom{N}{x_1, x_2, \dots, x_N} \left(\frac{1}{N}\right)^N$$

where,

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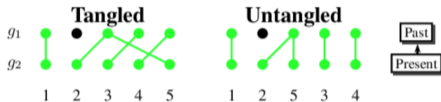


THE WRIGHT-FISHER GENEALOGY

- The W-F Genealogy tracks the parent-offspring relations in the W-F Model

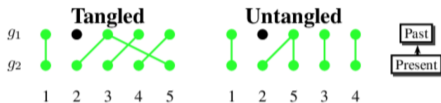
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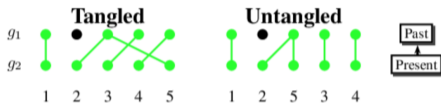
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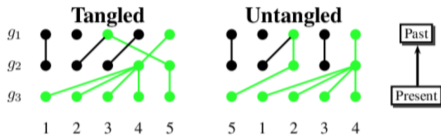
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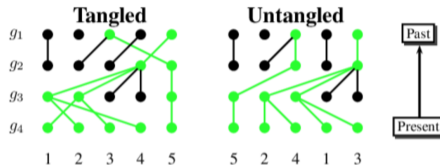
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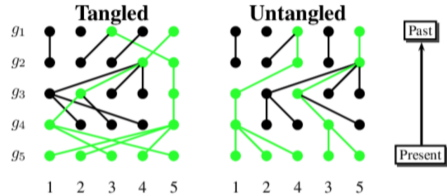
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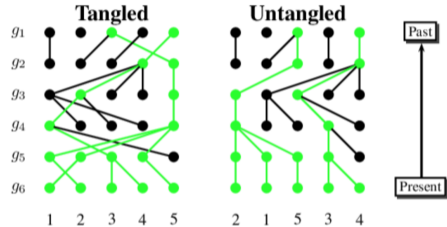
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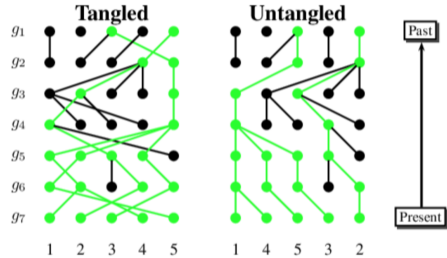
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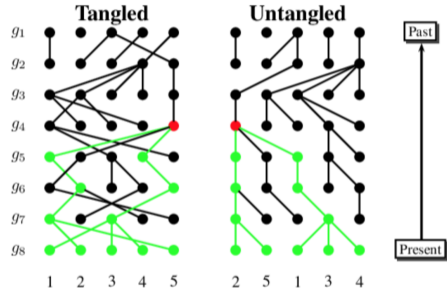
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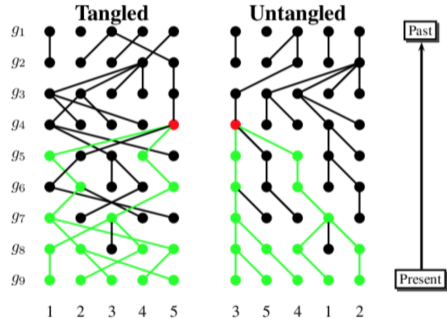
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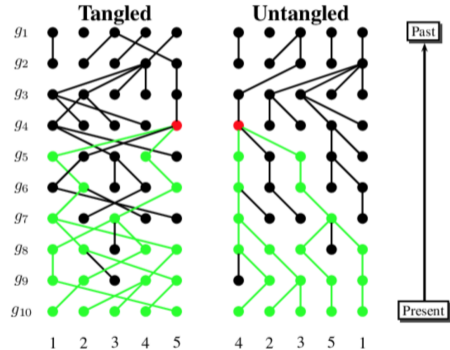
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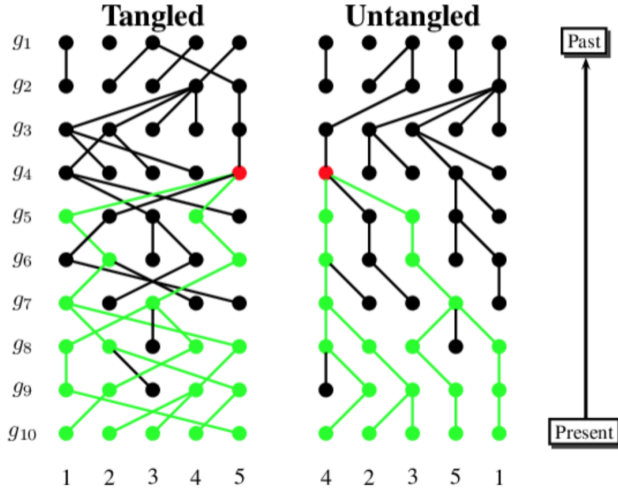


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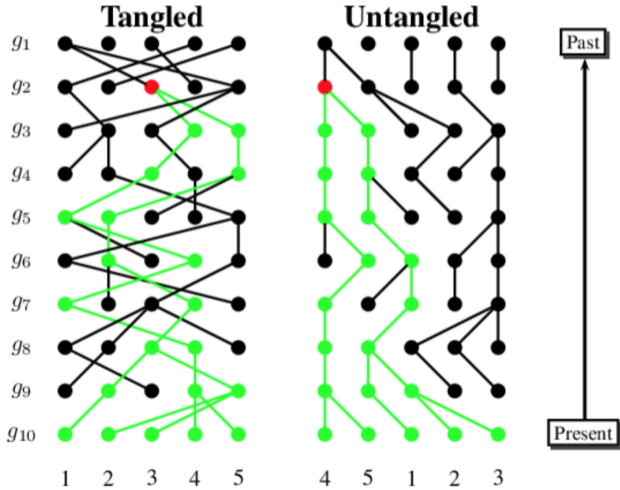
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- We want to study properties of the sample genealogy



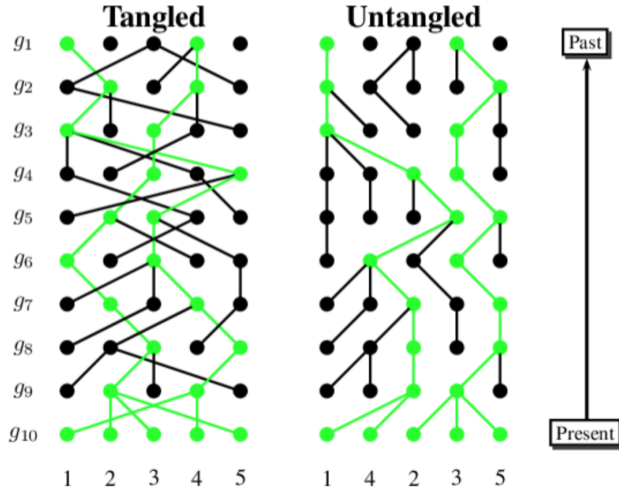
Genealogies are Random: 3 realisations when $N = 5$



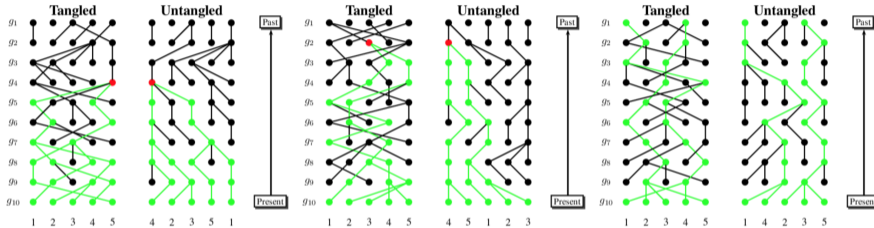
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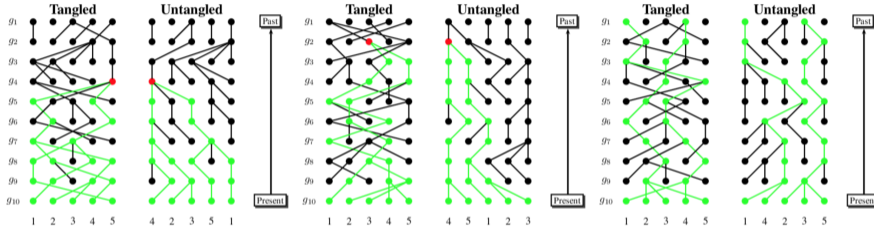


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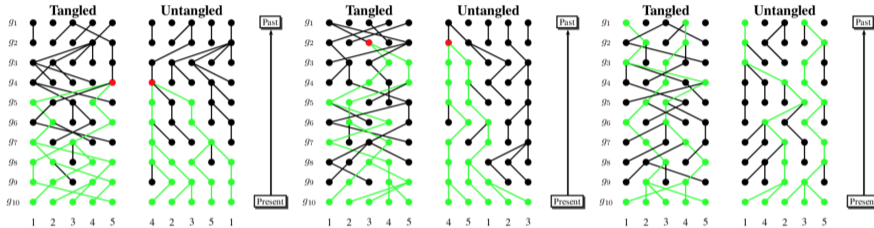
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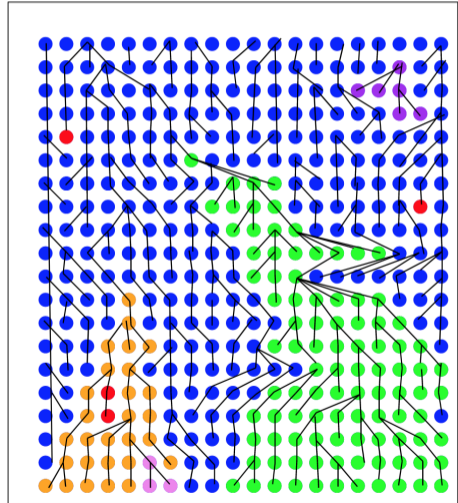
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- Moreover, the population size N is typically large ($N > 10^4$)
- We are often interested in the genealogy of a random sample of size n that is much smaller than N ($n \approx 30$) from the present generation
- Let us understand the idea of population and sample genealogies in a slightly larger population with $N = 20$ next

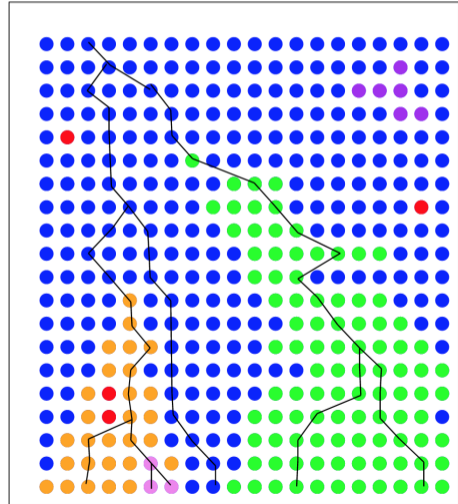
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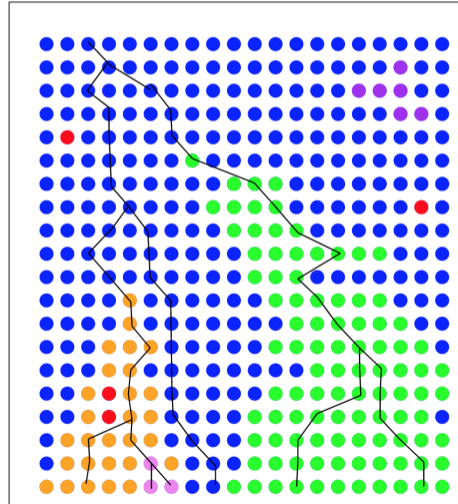
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- Note that the sample genealogy is a sub-graph of the population genealogy and merely comes from only tracking the parents, parents' parents, and so on of the six sampled individuals.
- Kingman's n -Coalescent Model, we study next, is a good approximation of the W-F sample genealogy and can shed light on some of its basic properties when n is smaller than N (Kingman, J. F. C. 1982. The coalescent. Stochastic Processes and Their Applications 13: 235-248.).



Mathematical Preliminaries 1 – a basic limit

Lemma

$$\lim_{N \rightarrow \infty} \left(1 - \frac{c}{N}\right)^{Nt} = e^{-ct}, \quad \text{for constants } c \text{ and } t > 0.$$

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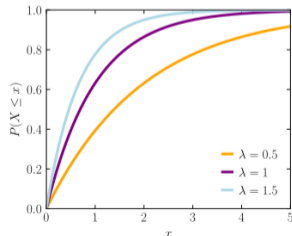
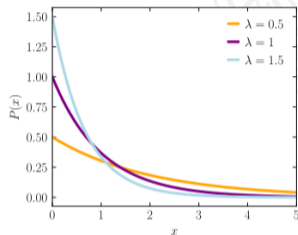
Mathematical Preliminaries 2 – Exponential(λ) R.V.

The probability density function (PDF) of Exponential(λ) R.V. with rate parameter $\lambda > 0$ is

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0, \\ 0 & x < 0. \end{cases}$$

And its cumulative distribution function (CDF) is

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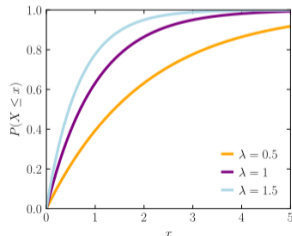
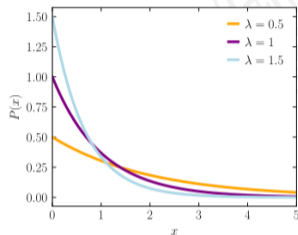
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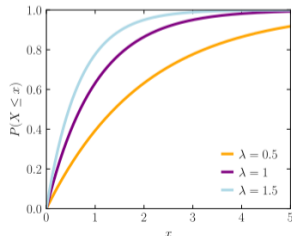
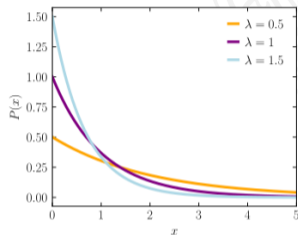
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Lemma (Competing Exponentials)

Let X_1, X_2, \dots, X_n be independent exponentially distributed random variables with rate parameters $\lambda_1, \lambda_2, \dots, \lambda_n$. Then $\min\{X_1, \dots, X_n\}$ is also exponentially distributed, with rate parameter $\lambda = \lambda_1 + \dots + \lambda_n$.





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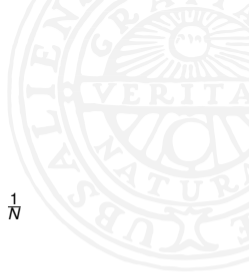
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- Let rescaled continuous time t be g in units of N generations. Then,
- $\Pr(2 \text{ individuals do not coalesce by rescaled time } t) = \left(1 - \frac{1}{N}\right)^{\lfloor Nt \rfloor} \xrightarrow{N \rightarrow \infty} e^{-t}$
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- Note that $1 - e^{-t}$ is the distribution function of the $\text{Exponential}(1)$, the Exponential random variable with rate parameter 1



Kingman's n -Coalescent: sample of size 2

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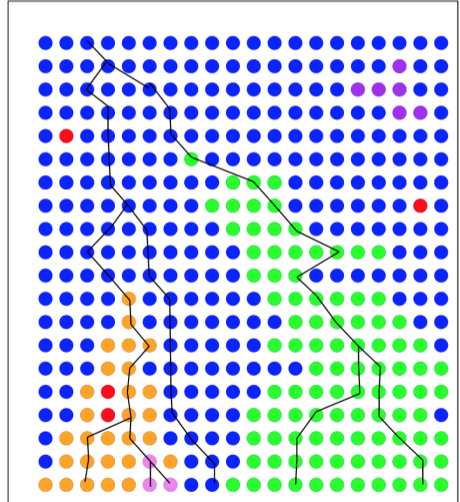


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- Thus the time for two individuals to coalesce is given by $\text{Exponential}(1)$ random variable with density function e^{-t}
- This is the main idea of Kingman's approximation for the coalescence time of a random sample of two individuals from a population whose size N gets large in the limit.

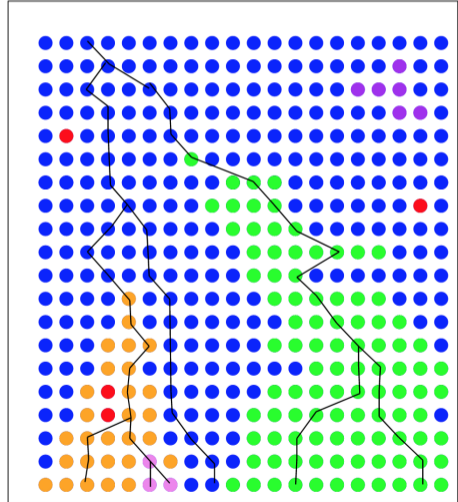
Kingman's n -Coalescent: sample of size $2 \leq i \leq n$

- Consider the genealogy of a random sample of size $n = 6$ from the present generation to fix ideas and let $2 \leq i \leq n$
- When there are i ancestral lineages of the sample that have not coalesced yet, there are $\binom{i}{2}$ pairs of lineages, each of which coalesce independently according to Exponential(1) random variable.



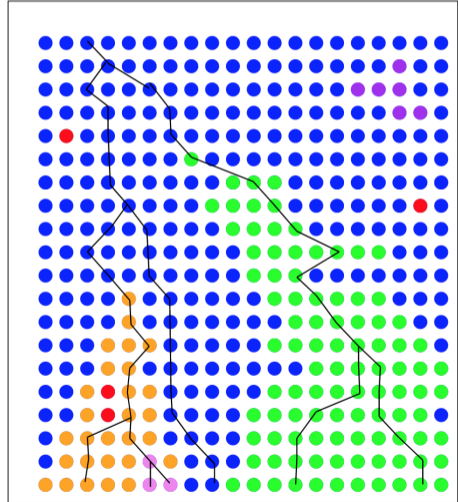
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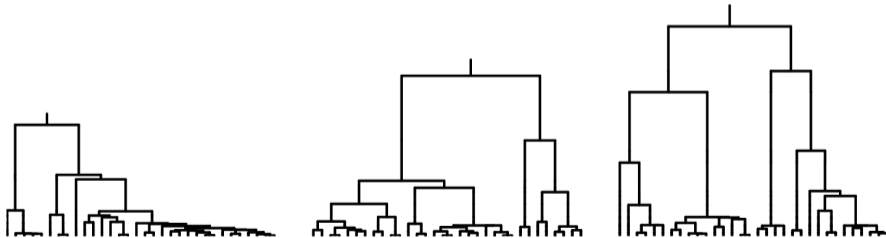
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- This already gives us a simulation algorithm of sample genealogies under Kingman's n -Coalescent Model.





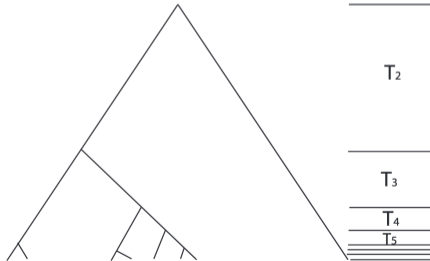
Kingman's n -Coalescent: Simulating Genealogies

- Draw a sample from $\text{Exponential}(\binom{n}{2})$ random variable.
- Pick two lineages at random to coalesce.
- Decrease n by 1
- If $n = 1$, stop; else, repeat these steps.



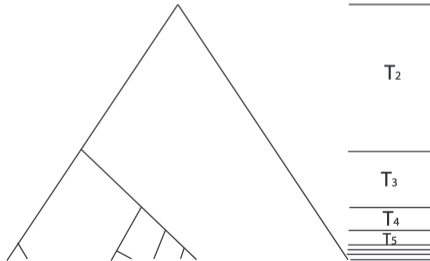
Kingman's n -Coalescent: Genealogical Times

- Let T_i be the time in the history of the sample when there were exactly i ancestral lineages



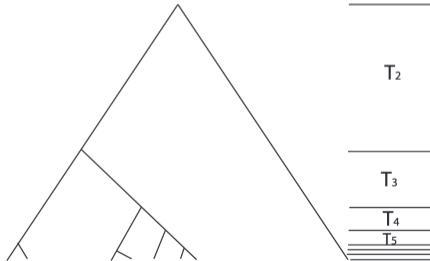
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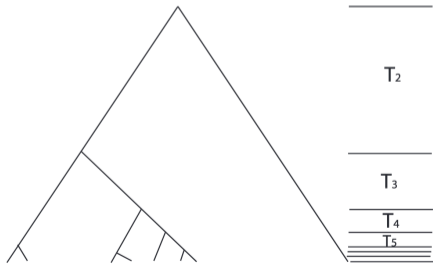
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Kingman's n -Coalescent: Genealogical Times

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- Total time represented by the genealogy = $T_{\text{TOT}} = \sum_{i=2}^n iT_i$
- We are interested in properties of T_{MRCA} and T_{TOT} .



Kingman's n -Coalescent: Means of T_{MRCA} and T_{TOT}

- The mean or expected time to the MRCA is

$$E(T_{\text{MRCA}}) = E\left(\sum_{i=2}^n T_i\right) = \sum_{i=2}^n E(T_i) = \sum_{i=2}^n \frac{2}{i(i-1)} = 2\left(1 - \frac{1}{n}\right)$$

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- The Variance as well as the probability densities of T_{MRCA} and T_{TOT} are also known.

Summary

Kingman's n -Coalescent model is a much simpler approximation to the sample genealogy of the Wright-Fisher model, provided we have a random sample of size n that is much smaller than the population size N , by rescaling time in units of N generations.

Extensions to Coalescent Models

All these extensions are possible and routinely used in genomic sciences today:

- Allow DNA sequences obtained from the sampled individuals to coalesce through the genealogies by also modeling DNA mutations to occur at given mutation rates
- Allow models to have demographic changes through time
- Allow sub-populations structure and migration between them
- Allow some DNA sequences to recombine and be trackable into both parents
- Allow samples from populations of different species
- etc.
- And, finally use the extended coalescent models for inference (testing and prediction) based on sampled DNA sequences

Coalescent Models with Extensions in Action

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Initial Upper Palaeolithic humans in Europe had recent Neanderthal ancestry

[Mateja Hajdinjak](#) , [Fabrizio Mafessoni](#), [Laurits Skov](#), [Benjamin Vernot](#), [Alexander Hübner](#), [Qiaomei Fu](#), [Elena Essel](#), [Sarah Nagel](#), [Birgit Nickel](#), [Julia Richter](#), [Oana Teodora Moldovan](#), [Silviu Constantin](#), [Elena Endarova](#), [Nikolay Zahariev](#), [Rosen Spasov](#), [Frido Welker](#), [Geoff M. Smith](#), [Virginie Sinet-Mathiot](#), [Lindsey Paskulin](#), [Helen Fewlass](#), [Sahra Talamo](#), [Zeljko Rezek](#), [Svoboda Sirakova](#), [Nikolay Sirakov](#), ... [Svante](#)

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Abstract

Modern humans appeared in Europe by at least 45,000 years ago^{1,2,3,4,5}, but the extent of their interactions with Neanderthals, who disappeared by about 40,000 years ago⁶, and their relationship to the broader expansion of modern humans outside Africa are poorly understood. Here we present genome-wide data from three individuals dated to between



Outlook: Current Research Perspectives

- Empirical Pedigree Processes
- Diploid Recombining Ancestries

Empirical Pedigree Processes

Inspired by Chatham Island Black Robins.



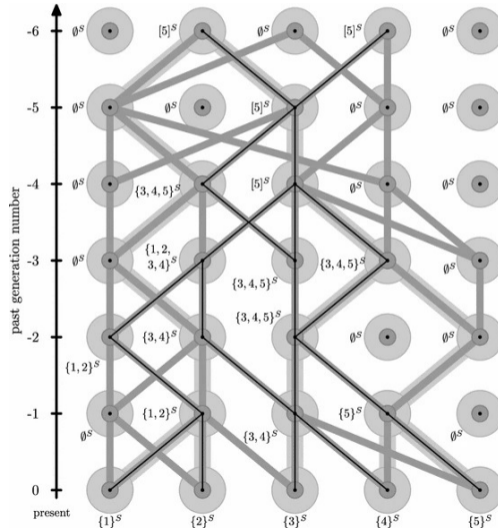
Empirical Pedigree Processes

Insights from field work in November 2011

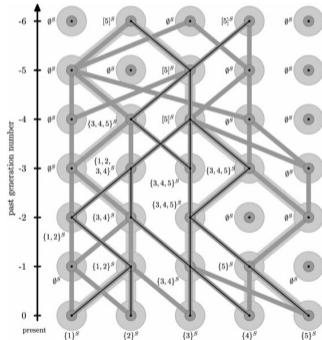


Inspired by Chatham Island Black Robins.

Diploid Recombining Ancestral Process



Diploid Recombining Ancestral Process – TMRCA



Theorem (Sainudiin-Thatte-Veber)

Let T_n denote the number of generations, counting back in time from the present, to an MRCA of all present-day individuals. Then for every $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}\{(1 - \varepsilon)C(r) \ln n \leq T_n \leq (1 + \varepsilon)C(r) \ln n\} = 1,$$

$$C(r) := \frac{1}{\ln(1+r)} - \frac{1}{\ln(1-r)}.$$



Want to learn more?

I recommend lectures by:

- Mark Beaumont https://gtpb.github.io/PGDH19/assets/coal_course/lecture1.pdf
- John Wakeley <https://www.youtube.com/watch?v=lo8641Pyszg>

Classical Introductory Textbook by:

- Coalescent Theory: An Introduction by Wakeley. <https://www.macmillanlearning.com/college/us/product/Coalescent-Theory/p/0974707759>